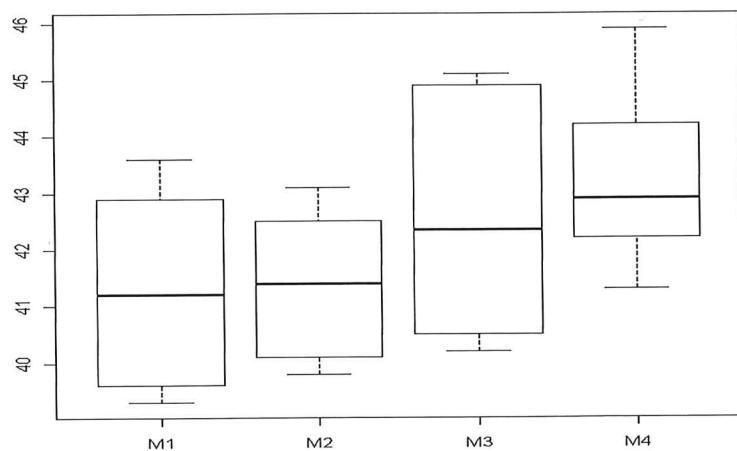


## The machine example

Four different machines  $M_1, M_2, M_3$  and  $M_4$  are being considered for the assembling of a particular product. It was decided that six different operators would be used in an experiment to compare the machines. The response is the amount of time required to assemble a product. The operation of the machines requires physical dexterity, and it was anticipated that there would be a difference among the operators in the speed with which they operated the machines. This difference causes undesired variability in the measured response which would be of interest to eliminate. Therefore, it was decided to perform the experiment as a randomized block design, with operators as blocks. The experiments will now be randomized within blocks. The observed data are given below:

Operators	Machines			
	$M_1$	$M_2$	$M_3$	$M_4$
1	42.5	39.8	40.2	41.3
2	39.3	40.1	40.5	42.2
3	39.6	40.5	41.3	43.5
4	39.9	42.3	43.4	44.2
5	42.9	42.5	44.9	45.9
6	43.6	43.1	45.1	42.3

## Boxplot maskiner



Let  $H_0: \alpha_1 = \alpha_2 = \dots = \alpha_k = 0$        $H_1$  at least one  $\alpha_i$  different from zero.

Under  $H_0$ :

$$\frac{SS_A}{\sigma^2} = \sum_{i=1}^k \sum_{j=1}^b (\bar{y}_{ij} - \bar{y}_{..})^2 \sim \chi^2_{(k-1)}$$

$$\frac{SS_E}{\sigma^2} = \sum_{i=1}^k \sum_{j=1}^b \frac{(\bar{y}_{ij} - \bar{y}_{..})^2}{\sigma^2} \sim \chi^2_{(k-1)(b-1)}$$

Det kan visast at  $\frac{SS_A}{\sigma^2} = \frac{\sum_{i=1}^k \sum_{j=1}^b (\bar{y}_{ij} - \bar{y}_{..} - \bar{y}_j + \bar{y}_{..})^2}{\sigma^2}$

er  $\chi^2_{((k-1)(b-1))}$  altså

such that  $F = \frac{\frac{SS_A}{\sigma^2(k-1)}}{\frac{SS_E}{\sigma^2(k-1)(b-1)}}$  er  $F$  fordelet med  $(k-1)$  og  $(k-1)(b-1)$  fridomsgrader.

Reject  $H_0$  if.  $F_{obs} \geq F_{\alpha, (k-1), (k-1)(b-1)}$

### Variansanalyse tabelle

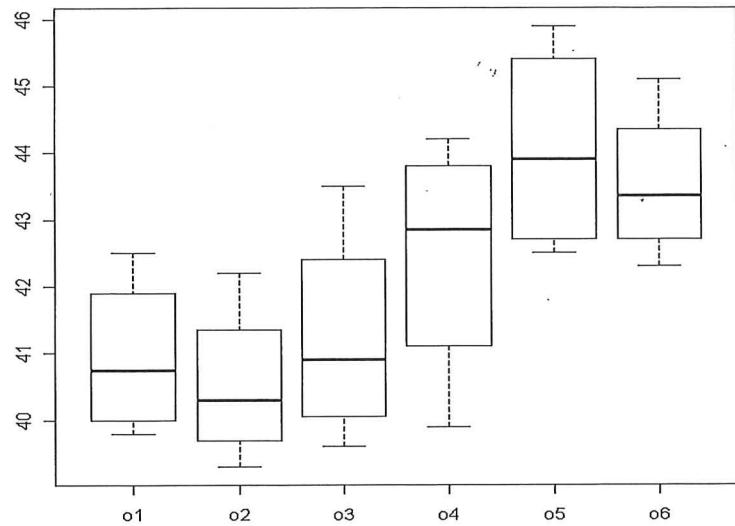
Kilder til variasjon	SS	DF	MS	F
Behandlinger	$SS_A = b \sum_{i=1}^k (\bar{y}_{..} - \bar{y}_{..})^2$	$k-1$	$SS_A / (k-1)$	$\frac{SS_A}{(k-1)} / \frac{SS_E}{(k-1)(b-1)}$
Blokker	$SS_B = k \sum_{j=1}^b (\bar{y}_{..j} - \bar{y}_{..})^2$	$b-1$	$SS_B / (b-1)$	
Feil	$SS_E = \sum_{i=1}^k \sum_{j=1}^b (\bar{y}_{ij} - \bar{y}_{..} - \bar{y}_{..j} + \bar{y}_{..})^2 / (k-1)(b-1)$		$SS_E / ((k-1)(b-1))$	
Totalt	$SS_T = \sum_{i=1}^k \sum_{j=1}^b (y_{ij} - \bar{y}_{..})^2$	$bk-1$		

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )	
Machine M1	41.3000	0.7413	55.72	<2e-16	***
Machine M2	41.3833	0.7413	55.83	<2e-16	***
Machine M3	42.5667	0.7413	57.42	<2e-16	***
Machine M4	43.2333	0.7413	58.33	<2e-16	***

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Machines	3	15.925	5.3082	1.6101	0.2186
Residuals	20	65.935	3.2968		

### Boxplot operators



Coefficients:

	Estimate	Std. Error	t value	Pr(> t )	
Operator o1	40.9500	0.7432	55.10	<2e-16	***
Operator o2	40.5250	0.7432	54.52	<2e-16	***
Operator o3	41.2250	0.7432	55.47	<2e-16	***
Operator o4	42.4500	0.7432	57.12	<2e-16	***
Operator o5	44.0500	0.7432	59.27	<2e-16	***
Operator o6	43.5250	0.7432	58.56	<2e-16	***

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Machine	3	15.925	5.3082	3.3388	0.047904 *
Operator	5	42.087	8.4174	5.2944	0.005328 **
Residuals	15	23.848	1.5899		

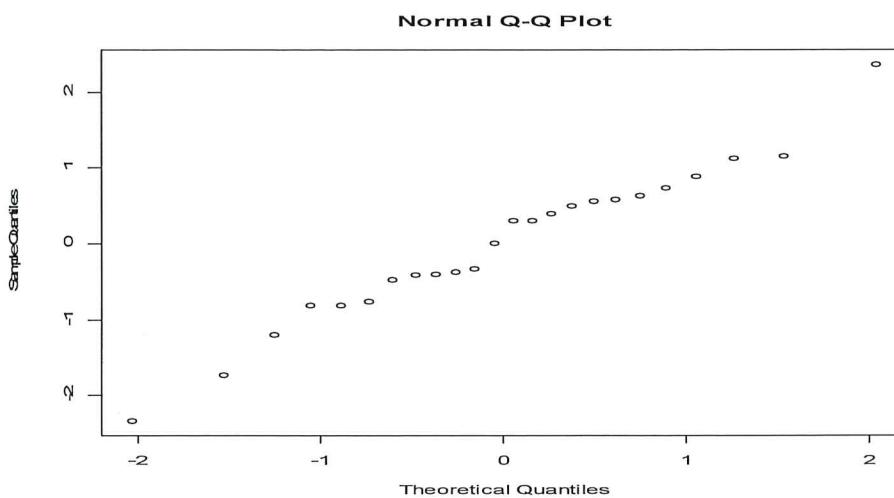
Suppose we neglected the blocking and analysed as a one-way ANOVA  
We have

$SS_E^{\text{one-way}} = SS_E^{\text{block}} + SS_B$  showing that by blocking  
the experiment we may extract unwanted variation from the  
data at the expense of losing some degrees of freedom

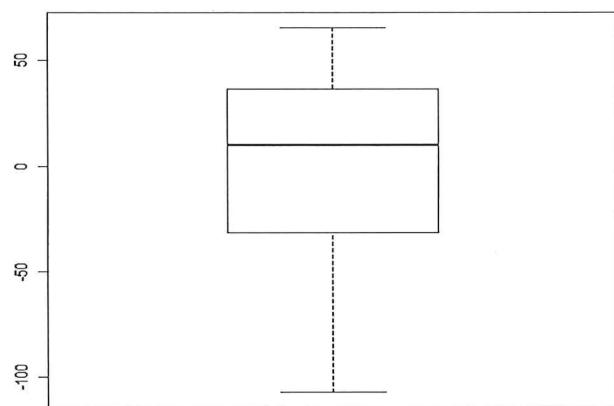
NB

The residuals are given by  $y_{ij} - \bar{y}_{i\cdot} - \bar{y}_{\cdot j} + \bar{y}_{\cdot\cdot}$   
 $= y_{ij} - (\underbrace{\bar{y}_{i\cdot} + \bar{y}_{\cdot j} - \bar{y}_{\cdot\cdot}}_{\hat{y}_{ij} \text{ (fitted model)}})$

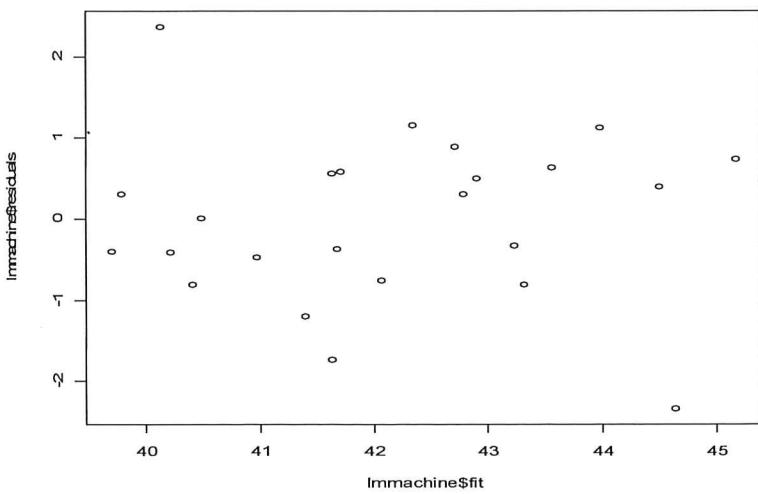
If residuals are suspected to be outliers they should  
be plotted against treatment and blocks.



## Boxplot residuals



## Residualplott



## Two-factor analysis of variance with replication

Situation: We have two factors A and B. A has  $a$  levels and B has  $b$  levels. By doing replications it is possible to study interactions between the factors. The experiment should be completely randomized, i.e. all the  $abn$  experiments, where  $n$  is the number of replicates should be performed in random order.

Model: 
$$Y_{ijk} = \mu + \alpha_i + \beta_j + \alpha\beta_{ij} + \varepsilon_{ijk} \quad \left\{ \begin{array}{l} N(0, \sigma^2) \text{ and independent} \\ i=1, 2, \dots, a, j=1, 2, \dots, b \\ k=1, 2, \dots, n \end{array} \right.$$

where  $\alpha_i$  is the effect of the  $i$ -th level of factor A,  $\beta_j$  is the effect of the  $j$ -th level of factor B and  $\alpha\beta_{ij}$  is the interaction between the  $i$ -th level of factor A and the  $j$ -th level of factor B.

$$\text{Also } \sum_{i=1}^a \alpha_i = \sum_{j=1}^b \beta_j = \sum_{i=1}^a \alpha\beta_{ij} = \sum_{j=1}^b \alpha\beta_{ij} = 0$$

The following hypothesis are of interest.

1.  $H_0: \alpha_1 = \alpha_2 = \dots = \alpha_a = 0 \quad H_1: \text{at least one is different from 0.}$
2.  $H_0: \beta_1 = \beta_2 = \dots = \beta_b = 0 \quad H_1: \text{at least one is different from 0.}$
3.  $H_0: \alpha\beta_{11} = \alpha\beta_{12} = \dots = \alpha\beta_{ab} = 0 \quad H_1: \text{at least one is different from 0.}$

## Partitioning variation

We have  $y_{ijk} = \bar{y}_{ij} + y_{ijk} - \bar{y}_{ij}$ ,  
 and  $\bar{y}_{ij}$  may be partitioned the same way as  $y_{ij}$   
 for the unreplicated two-way ANOVA (or randomized  
 complete block design). Therefore,

$$\bar{y}_{ij} = \bar{y}_{..} + \underbrace{\bar{y}_{i..} - \bar{y}_{..}}_{\hat{\alpha}_i} + \underbrace{\bar{y}_{j..} - \bar{y}_{..}}_{\hat{\beta}_j} + \underbrace{\bar{y}_{ij..} - \bar{y}_{i..} - \bar{y}_{j..} + \bar{y}_{..}}_{\hat{\alpha}\hat{\beta}_{ij}}$$

and we get,

$$y_{ijk} = \bar{y}_{..} + \underbrace{\bar{y}_{i..} - \bar{y}_{..}}_{\hat{\alpha}_i} + \underbrace{\bar{y}_{j..} - \bar{y}_{..}}_{\hat{\beta}_j} + \underbrace{\bar{y}_{ij..} - \bar{y}_{i..} - \bar{y}_{j..} + \bar{y}_{..}}_{\hat{\alpha}\hat{\beta}_{ij}} + y_{ijk} - \bar{y}_{ij}$$

Also  $y_{ijk} - \bar{y}_{..} = \bar{y}_{ij} - \bar{y}_{..} + y_{ijk} - \bar{y}_{ij}$ .

Since  $\sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^m (\bar{y}_{ij} - \bar{y}_{..})(y_{ijk} - \bar{y}_{ij}) = 0$

We obtain

$$\begin{aligned} \sum_{l=1}^a \sum_{j=1}^b \sum_{k=1}^m (y_{ijk} - \bar{y}_{..})^2 &= \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^m (\bar{y}_{ij} - \bar{y}_{..})^2 + \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^m (y_{ijk} - \bar{y}_{ij})^2 \\ &= \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^m (\bar{y}_{i..} - \bar{y}_{..})^2 + \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^m (\bar{y}_{j..} - \bar{y}_{..})^2 + \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^m (\bar{y}_{ij..} - \bar{y}_{i..} - \bar{y}_{j..} + \bar{y}_{..})^2 \\ &\quad + \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^m (y_{ijk} - \bar{y}_{ij})^2 \end{aligned}$$

or,  $SS_T = SS_A + SS_B + SS_{AB} + SS_E$